Influence of Weight Reduction on Blood Pressure
A Meta-Analysis of Randomized Controlled Trials

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Abstract—Increased body weight is a strong risk factor for hypertension. A meta-analysis of randomized controlled trials was performed to estimate the effect of weight reduction on blood pressure overall and in population subgroups. Twenty-five randomized, controlled trials (comprising 34 strata) published between 1966 and 2002 with a total of 4874 participants were included. A random-effects model was used to account for heterogeneity among trials. A net weight reduction of \(-5.1\) kg (95% confidence interval [CI], \(-6.03\) to \(-4.25\)) by means of energy restriction, increased physical activity, or both reduced systolic blood pressure by \(-4.44\) mm Hg (95% CI, \(-5.93\) to \(-2.95\)) and diastolic blood pressure by \(-3.57\) mm Hg (95% CI, \(-4.88\) to \(-2.25\)). Blood pressure reductions were \(-1.05\) mm Hg (95% CI, \(-1.43\) to \(-0.66\)) systolic and \(-0.92\) mm Hg (95% CI, \(-1.28\) to \(-0.55\)) diastolic when expressed per kilogram of weight loss.

As expected, significantly larger blood pressure reductions were observed in populations with an average weight loss >5 kg than in populations with less weight loss, both for systolic (\(-6.63\) mm Hg [95% CI, \(-8.43\) to \(-4.82\)] vs \(-2.70\) mm Hg [95% CI, \(-4.59\) to \(-0.81\)]) and diastolic (\(-5.12\) mm Hg [95% CI, \(-6.48\) to \(-3.75\)] vs \(-2.01\) mm Hg [95% CI, \(-3.47\) to \(-0.54\)]) blood pressure. The effect on diastolic blood pressure was significantly larger in populations taking antihypertensive drugs than in untreated populations (\(-5.31\) mm Hg [95% CI, \(-6.64\) to \(-3.99\)] vs \(-2.91\) mm Hg [95% CI, \(-3.66\) to \(-2.16\)]). This meta-analysis clearly shows that weight loss is important for the prevention and treatment of hypertension. (Hypertension. 2003;42:878-884.)

Key Words: obesity ■ body weight ■ blood pressure ■ hypertension, obesity ■ meta-analysis ■ clinical trials

Hypertension is an established risk factor for cardiovascular diseases and is common among obese subjects.1 In Western societies, the prevalence of obesity is increasing, which is likely to increase the burden of hypertension and consequently, of cardiovascular mortality.2 From 1960 to 2000, the prevalence of obesity (body mass index [BMI] ≥30 kg/m²) for American adults aged 20 to 74 years increased from 13.4% to 30.9%. In 2000, the prevalence of overweight (BMI ≥25 kg/m²) for American adults was 64.5%.3 Weight loss has been proposed as an effective, nonpharmacologic means for the primary prevention of hypertension.4 An early meta-analysis by Staessen et al5 in 1988 showed a reduction in systolic blood pressure (SBP) and diastolic blood pressure (DBP) of \(-2.4\) and \(-1.5\) mm Hg per kilogram weight loss, respectively.

Targeted weight loss interventions in population subgroups might be more effective for the prevention of hypertension than a general-population approach because of interindividual differences in BP and body weight. In the present meta-analysis of 25 randomized controlled trials (RCTs), both the overall and subgroup effects of body weight reduction on BP were assessed. This stratified meta-analysis could strengthen the scientific basis for weight control as an important ingredient of hypertension control programs in the community and in individual patients.

Methods

Selection of Studies
Systematic literature searches in electronic databases (Cab Abstract, Current Contents, FSTA, and MEDLINE Advanced) were conducted for RCTs of weight reduction and BP by using the search phrase (overweight OR weight reduction OR obes* OR weight OR diet*) AND (hypertension OR blood pressure) AND (trial OR intervention OR random* OR study) as title words or MeSH terms. Article reference lists were examined for additional articles.

Trials that fulfilled the following criteria were eligible for meta-analysis: (1) an RCT design, (2) published after 1966, (3) conducted in humans, (4) written in English, and (5) nonpharmacologic reduction of body weight. A total of 97 articles was identified, of which 72 were excluded on the basis of the predefined criteria (literature list of excluded trials is available from the authors); i.e., they did not fulfill the inclusion criteria (n=21), the intervention period was <8 weeks (n=2), there were co-interventions from which the effect of weight loss could not be separated (n=28), inappropriate or missing BP data (n=5), study objective was not weight reduction (n=4), or overlapping data with other publication(s) (n=12). Figure 1 shows the number of studies that were identified and excluded at different stages of the selection process. A total of 25 trials6–30 proved eligible for this meta-analysis.
Data Abstraction

Original articles were retrieved for data abstraction. Two investigators (J.E.N., B.E.S.) independently collected data on population characteristics, study design, and treatment effect on a standard form. For trials in which BP and body weight measurements were recorded at several points in time,\textsuperscript{6,14} data were abstracted for the intervention period during which the maximal BP effect was achieved. If intermediate measurements were not reported, then the total intervention period was used.\textsuperscript{15,30} Different types of weight intervention could have been examined within the same trial, ie, energy restriction vs increased physical activity vs the combined intervention. When this was the case, these interventions were analyzed as separate strata. For 2 trials, the authors were contacted to obtain missing data.\textsuperscript{15,25} For 1 trial, in which the age of the trial population was not reported, the mean age of other trial populations combined (ie, 46 years) was imputed in the data analysis.\textsuperscript{28}

Quantitative Data Synthesis

Changes in BP from baseline in the control group were subtracted from changes in the intervention group to obtain the net BP effect of weight reduction. SEs of the treatment effect were also abstracted. If not reported, SEs were derived from SDs or SEs for BP effects within groups, confidence intervals (CIs), test statistics, or probability values. For estimations of SE, a correlation of 0.5 between the variances at baseline and at follow-up within each group was assumed, according to Follmann et al,\textsuperscript{31} as follows:

$$ SE = \sqrt{\left( \frac{1}{SE_{\text{baseline}}^2} + \frac{1}{SE_{\text{follow-up}}^2} - 2 \times 0.5 \times \frac{SE_{\text{baseline}}}{SE_{\text{follow-up}}} \right)} $$

In one article,\textsuperscript{26} variance measures were not presented, so SE was predicted from other trials by means of linear regression by using the trial sample size as the independent variable.

A random-effects model was chosen for meta-analysis to account for heterogeneity in BP effect among trials. An advanced statistical technique for meta-analysis of continuous outcomes was applied that takes into account both within- and between-study variation (SAS PROC MIXED).\textsuperscript{32} The weight factor by which BP effects were weighted was $1/(SE^2 + \tau^2)$, where $\tau$ is the between-study variation.

To examine the effect modification of BP response to weight loss, meta-analyses were performed in predefined strata of mean age ($\leq 45$ vs $>45$ years), gender ($<50\%$ vs $\geq 50\%$ females), type of intervention (energy restriction vs increased physical activity vs combined intervention), race (white vs black vs Asian), initial BMI ($<30$ vs $\geq 30$ kg/m$^2$), antihypertensive treatment (no vs yes; coded as "yes" if 1 or more subjects were being treated with antihypertensive drugs), weight loss ($\leq 5$ vs $>5$ kg), and baseline hypertensive status (yes vs no; based on initial BP level only, ie, $<140/90$ vs $\geq 140/90$ mm Hg). Stratified meta-analyses were repeated with a multivariate model to adjust for potential confounders, ie, mean age, gender (percent females), initial BP, change in body weight of trial populations, and duration of intervention.

A funnel plot of SBP effect against the weight factor was visually examined for potential publication bias (Figure 2). The funnel plot was roughly symmetrical about the mean-effect size line, except for the outlier study by Reisin et al.\textsuperscript{27} For this reason, meta-analyses were also performed after exclusion of this study.

Results

Overview of Trials

The population and study characteristics of trials of weight reduction and BP are presented in Table 1. The analysis was based on 25 parallel trials published between 1978 and 2002, comprising 34 strata with a total of 4874 subjects. The duration of intervention until the maximal BP effect (or, if not reported the overall BP effect) was achieved varied between 8 and 260 weeks. Most studies (82%) included both men and women. The mean age of trial populations ranged from 37 to 66 years. Half of the populations were hypertensive (on the basis of initial BP cutoff level of $140/90$ mm Hg), and 24% of the populations were taking antihypertensive medication. Mean total duration was 66.6 weeks, and mean duration until the maximal BP effect was achieved was 35.3 weeks. The overall percentage of subjects in RCTs who dropped out after randomization was small (4.8%).
Mean initial body weight and BMI were 88.3 kg and 30.7 kg/m², respectively. Mean net change in body weight of the population was −5.8% of initial body weight, i.e., −5.1 kg (95% CI, −6.03 to −4.25), ranging from −0.6 to −11.9 kg. Average weight reductions in the energy restriction, increased physical activity, and combined intervention subgroups were −6.7 kg (95% CI, −8.27 to −5.11), −3.1 kg (95% CI, −4.54 to −1.75), and −6.2 kg (95% CI, −7.87 to −4.55), respectively.

Changes in BP
Forest plots for net changes in SBP and DBP due to weight reduction, with 95% CIs, are presented in Figure 3. An average net weight reduction of −5.1 kg by means of energy restriction, increased physical activity, or the combined intervention was associated with a significant reduction in SBP of −4.44 mm Hg (95% CI, −5.93 to −2.95) and a significant reduction in DBP of −3.57 mm Hg (95% CI, −4.88 to −2.25). When expressed per kilogram of weight loss, reductions in SBP and DBP were −1.05 mm Hg (95% CI, −1.43 to −0.67) and −0.67 mm Hg (95% CI, −0.83 to −0.51), respectively.
2. In the stratified meta-analysis, larger BP reductions were observed in populations with an average weight loss >5 kg than in populations with less weight loss, both for SBP (−6.24 mm Hg [95% CI, −8.06 to −4.41] vs −2.44 mm Hg [95% CI, −4.38 to −0.49]) and DBP (−4.97 mm Hg [95% CI, −6.62 to −3.31] vs −1.97 mm Hg [95% CI, −3.71 to −0.21]). BP reductions were also larger in populations who were taking antihypertensive medication than in untreated populations, both for SBP (−7.00 mm Hg [95% CI, −10.02 to −3.98] vs −3.77 mm Hg [95% CI, −5.33 to −2.22]) and DBP (−5.49 mm Hg [95% CI, −8.06 to −2.93] vs −2.97 mm Hg [95% CI, −4.39 to −1.55]). When subgroups were compared on the basis of initial BP level (<140/90 mm Hg vs ≥140/90 mm Hg), there was no difference in SBP response, but reductions in DBP were approximately twice as large in hypertensives, although this difference was not statistically significant (−4.92 [95% CI, −6.73 to −3.12] vs −2.35 [95% CI, −4.05 to −0.65]). Weight loss caused larger BP reductions in Asian than in white or black populations, both for SBP and DBP, but findings must be interpreted with caution because of limited statistical power. Age, initial BMI, gender (percent females), and type of intervention did not significantly influence SBP and DBP response to weight loss in unadjusted analyses (Table 2).

Excluding the study of Reisin et al.27 yielded a smaller contrast in BP response between treated and untreated populations, ie, −5.87 mm Hg (95% CI, −8.79 to −2.94) versus −3.79 mm Hg [95% CI, −5.25 to −2.32] for SBP and −3.90 mm Hg (95% CI, −6.10 to −1.70) versus −2.90 mm Hg (95% CI, −4.06 to −1.74) for DBP.

Findings from multivariate stratified analyses were consistent with univariate results (Table 2), except for strongly increased BP estimates for intervention with physical activity (−4.74 mm Hg [95% CI, −7.60 to −1.88] for SBP and −4.65 mm Hg [95% CI, −6.84 to −2.45] for DBP), which resulted from correction for the amount of weight reduction.

### Discussion

This meta-analysis of 25 RCTs comprising 4874 participants from different ethnic populations showed a BP reduction of −4.4/−3.6 mm Hg for an ~5-kg weight loss by means of energy restriction, physical activity, or both. Larger BP reductions were achieved in populations that included subjects taking antihypertensive drugs.

Our meta-analysis was based on trials with an RCT design, which have high internal validity. For trials in which BP measurements were repeatedly recorded during follow-up, data were abstracted at the time point when the maximal effect of weight reduction on BP was achieved (32% of the strata). In most of these RCTs, the maximal effect was reached before the end of the trial. A likely explanation is lack of compliance during long-term intervention.

A previous meta-analysis of 12 trials by Staessen et al.5 published in 1988, showed a BP reduction of −2.4/−1.5 mm Hg per kilogram weight loss, which is substantially larger than our estimate (−1.1/−0.9 mm Hg per kg). However, half of the studies in the meta-analysis by Staessen et al.5 were not randomized, and only 2 trials overlapped with our study, which might explain this discrepancy. The results of our subgroup analysis by hypertensive status are not consis-
A meta-analysis of aerobic exercise and BP by Whelton et al.\textsuperscript{15} showed that BP was significantly reduced even in trials in which overall weight loss was minimal. This suggests that exercise reduces BP independently of changes in body weight. In our multivariate analysis, which was standardized for the amount of weight loss, the effect on DBP was larger when body weight was reduced by physical activity compared with energy restriction. This confirms the idea that physical activity also reduces BP by mechanisms unrelated to weight loss.

Although the exact mechanism of the relation between hypertension and obesity and the effect of weight loss on BP is unknown, there are several plausible biologic pathways.\textsuperscript{1} The renin-angiotensin-aldosterone system is overactivated in obese subjects, and renin activity and aldosterone concentrations are higher than in lean subjects.\textsuperscript{36–38} Furthermore, activity of the sympathetic nervous system is increased in

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**TABLE 2. Changes in SBP and DBP in 25 RCTs of Weight Reduction and BP, Overall and in Subgroups**

<table>
<thead>
<tr>
<th>Stratum</th>
<th>No. of Strata*</th>
<th>SBP, mm Hg\textsuperscript{†}</th>
<th>DBP, mm Hg\textsuperscript{†}</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>Unadjusted</td>
<td>Adjusted\textsuperscript{‡}</td>
</tr>
<tr>
<td>Overall</td>
<td>34</td>
<td>-4.44 (-5.93; -2.95)</td>
<td>-4.78 (-5.76; -3.80)</td>
</tr>
<tr>
<td>Age ≤45 years</td>
<td>15</td>
<td>-4.19 (-6.19; -2.20)</td>
<td>-4.74 (-6.35; -3.12)</td>
</tr>
<tr>
<td>Age &gt;45 years</td>
<td>19</td>
<td>-4.74 (-6.85; -2.52)</td>
<td>-4.80 (-6.48; -3.13)</td>
</tr>
<tr>
<td>Gender &lt;50% females</td>
<td>21</td>
<td>-4.75 (-6.54; -2.97)</td>
<td>-5.05 (-6.10; -3.99)</td>
</tr>
<tr>
<td>Gender ≥50% females</td>
<td>13</td>
<td>-3.74 (-6.40; -1.07)</td>
<td>-3.91 (-5.69; -2.13)</td>
</tr>
<tr>
<td>Hypertension§</td>
<td>No</td>
<td>17</td>
<td>-4.08 (-6.01; -2.16)</td>
</tr>
<tr>
<td>Hypertension§</td>
<td>Yes</td>
<td>17</td>
<td>-4.95 (-7.25; -2.64)</td>
</tr>
<tr>
<td>Race∥</td>
<td>White</td>
<td>14</td>
<td>-3.19 (-4.78; -1.59)</td>
</tr>
<tr>
<td>Race∥</td>
<td>Black</td>
<td>4</td>
<td>-4.67 (-8.86; -0.49)</td>
</tr>
<tr>
<td>Race∥</td>
<td>Asian</td>
<td>4</td>
<td>-8.77 (-11.91; -5.64)</td>
</tr>
<tr>
<td>Intervention Energy restriction</td>
<td>19</td>
<td>-4.93 (-6.84; -3.02)</td>
<td>-4.33 (-5.70; -2.97)</td>
</tr>
<tr>
<td>Intervention Physical activity</td>
<td>8</td>
<td>-1.73 (-5.14; 1.69)</td>
<td>-4.74 (-7.60; -1.88)</td>
</tr>
<tr>
<td>Intervention Combined intervention</td>
<td>7</td>
<td>-5.15 (-7.78; -2.51)</td>
<td>-5.66 (-7.52; -3.81)</td>
</tr>
<tr>
<td>Initial BMI &lt;30 kg/m\textsuperscript{2}</td>
<td>15</td>
<td>-4.14 (-4.95; -3.33)</td>
<td>-4.59 (-5.70; -3.49)</td>
</tr>
<tr>
<td>Initial BMI ≥30 kg/m\textsuperscript{2}</td>
<td>13</td>
<td>-4.09 (-4.87; -3.31)</td>
<td>-4.05 (-5.06; -3.05)</td>
</tr>
<tr>
<td>Weight reduction ≤5 kg</td>
<td>16</td>
<td>-2.44 (-4.38; -0.49)</td>
<td>-2.70 (-4.59; -0.81)</td>
</tr>
<tr>
<td>Weight reduction &gt;5 kg</td>
<td>18</td>
<td>-6.24 (-8.06; -4.41)</td>
<td>-6.63 (-8.43; -4.82)</td>
</tr>
<tr>
<td>Antihypertensive drugs¶</td>
<td>No</td>
<td>26</td>
<td>-3.77 (-5.33; -2.22)</td>
</tr>
<tr>
<td>Antihypertensive drugs¶</td>
<td>Yes</td>
<td>8</td>
<td>-7.00 (-10.02; -3.98)</td>
</tr>
</tbody>
</table>

*No. of strata in subgroup analyses may not total 34 because of missing data.
†BP changes were obtained from a random-effects model, with 95% CI end points in parentheses.
‡Adjusted for age, gender (% females), initial BP, change in body weight, and duration.
§Based on initial BP level (<140/90 mm Hg vs ≥140 mm Hg), not on use of antihypertensive medication.
∥Races other than black, white, and Asian were not considered in subgroup analysis because of small numbers; power for multivariate analysis in strata of race was insufficient.
¶Coded “yes” if trial population included one or more subjects on antihypertensive drug treatment.

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hypertensive, obese subjects, which could induce obesity-related renal effects. Alternatively, there might be inhibition of the natriuretic peptides system, of which the functional effects are vasodilatation and natriuresis. Decreased insulin sensitivity and hyperinsulinemia as part of the metabolic syndrome might also form an essential link between obesity and hypertension, although this interrelation is still not fully understood.

Additional studies are needed to investigate the long-term effects of weight reduction on BP. It is also important to examine whether weight loss has a persistent beneficial effect on cardiovascular risk, even after regaining weight. Data on weight reduction and BP in children and adolescents are scanty. Trials in this field are clearly warranted because of the increasing prevalence of obesity and hypertension at younger ages, which is becoming a major medical problem and a threat to public health.

In conclusion, the results from this meta-analysis provide unequivocal evidence that weight loss makes an important contribution to the treatment of hypertension, especially in subjects taking antihypertensive medication. Prevention of weight gain is likely to have a large impact on the burden of hypertension and consequently, cardiovascular diseases in the general population.

References


